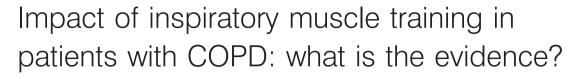
REVIEW



R. Gosselink*', J. De Vos*', S.P. van den Heuvel[§], J. Segers*', M. Decramer*', and G. Kwakkel⁺

ABSTRACT: A meta-analysis including 32 randomised controlled trials on the effects of inspiratory muscle training (IMT) in chronic obstructive pulmonary disease (COPD) patients was performed. Overall and subgroup analyses with respect to training modality (strength or endurance training, added to general exercise training) and patient characteristics were performed. Significant improvements were found in maximal inspiratory muscle strength ($P_{I,max}$; +13 cmH₂O), endurance time (+261 s), 6- or 12-min walking distance (+32 and +85 m respectively) and quality of life (+3.8 units). Dyspnoea was significantly reduced (Borg score -0.9 point; Transitional Dyspnoea Index +2.8 units). Endurance exercise capacity tended to improve, while no effects on maximal exercise capacity were found. Respiratory muscle endurance training revealed no significant effect on $P_{I,max}$, functional exercise capacity and dyspnoea. IMT added to a general exercise programme improved $P_{I,max}$ significantly, while functional exercise capacity tended to increase in patients with inspiratory muscle weakness ($P_{I,max} < 60 \text{ cmH}_2O$).

IMT improves inspiratory muscle strength and endurance, functional exercise capacity, dyspnoea and quality of life. Inspiratory muscle endurance training was shown to be less effective than respiratory muscle strength training. In patients with inspiratory muscle weakness, the addition of IMT to a general exercise training program improved *P*_{I,max} and tended to improve exercise performance.

KEYWORDS: Meta-analysis, muscle training, respiratory muscles, systematic review

espiratory muscle weakness is observed in chronic obstructive pulmonary disease (COPD) patients [1, 2] and contributes to hypercapnia [3], dyspnoea [4, 5], nocturnal oxygen desaturation [6] and reduced walking distance [7]. During exercise it has been shown that diaphragm work is increased in COPD [8] and COPD patients use a larger proportion of the maximal inspiratory pressure (PI,max) than healthy subjects [9]. This pattern of breathing is closely related to the dyspnoea sensation during exercise [9] and might potentially induce respiratory muscle fatigue. However, diaphragmatic fatigue was not demonstrated after exhaustive exercise [10]. Studies in patients with COPD have shown natural adaptations of the diaphragm to greater oxidative capacity and resistance to fatigue [11-13]. The abovementioned considerations gave conflicting

arguments to the rationale of respiratory muscle training in COPD. Current guidelines [14, 15] and meta-analyses [16-20] are not undisputedly positive on the application of inspiratory muscle training (IMT). From meta-analyses it is clear that IMT increases inspiratory muscle strength and endurance, and decreases dyspnoea. However, exercise performance and quality of life did not improve significantly [18, 20, 21]. The addition of IMT to a general exercise programme did not improve exercise performance [20]. Furthermore, differences in effects of resistance and endurance training were never analysed, while patient characteristics relevant for favourable effects of IMT have not been identified so far. Finally, many randomised controlled trials have been published since our previous meta-analysis, offering more input for more extensive analysis. The aim of the AFFILIATIONS

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European Respiratory Journal Print ISSN 0903-1936 Online ISSN 1399-3003 present meta-analysis is to: 1) investigate the effects of IMT as stand-alone therapy or added to general exercise training; 2) identify patient characteristics associated with favourable effects of IMT; and 3) identify the most appropriate training modality in terms of strength or endurance training for IMT.

METHODS

A detailed description of the methods section is available in the online supplement.

Study identification and selection

We performed an update and extension of the meta-analysis from our group [21]. For the current meta-analysis, an additional search up to May 2009 was performed independently by two reviewers (J. De Vos and S.P. van den Heuvel). Studies had to satisfy the following criteria to be included for further analysis: 1) true experimentation (i.e. randomised or quasirandomised controlled trial); 2) COPD patients with pulmonary function tests; 3) IMT at an intensity of $\ge 30\%$ PLmax [22] or respiratory muscle endurance training in a controlled manner; 4) outcomes should have been described in terms of inspiratory muscle strength (PI,max), inspiratory muscle endurance, dyspnoea rating, 6- or 12-min walking distance (6MWD or 12MWD) and/or health-related quality of life (HRQoL). In contrast to all previous meta-analyses, no language restrictions were used. The methodological quality of each included study was scored using a modification of the framework for methodological quality as previously described [16, 21].

Meta-analysis

The effect size gi (Hedges' g) of each individual study was calculated by the difference between the means of the experimental and the control groups before and after the intervention divided by the average population standard deviation (SDi). If data were not in a form suitable for quantitative pooling, we contacted trial authors for additional information. The analyses were initially performed for the total set of included studies. Subsequently, effectiveness was compared between inspiratory muscle strength (threshold or resistive) and endurance training and the addition of IMT to general exercise reconditioning (GER) was analysed. The homogeneity test Cochrane O-statistic (O-test) of each set of effect sizes was calculated to determine whether studies shared a common effect size of which the variance could be explained by sampling error alone. A minimum of three studies was set for pooling in order to perform a quantitative analysis of the studies. However, the Q-test is known to be poor at detecting true heterogeneity among studies as significant [23], in particular when a small number of studies is included and the power of the test is low. Therefore, the I²-statistic was calculated to determine heterogeneity and consider further sensitivity analysis. Heterogeneity was defined if I² was beyond 50% [23].

On the basis of existing literature, we hypothesised the following potential effect modifiers: 1) patient characteristics ($P_{I,max} \leq 60 \text{ or } > 60 \text{ cmH}_2\text{O}$, forced expiratory volume in 1 s, and arterial oxygen (P_{a,O_2}) and carbon dioxide (P_{a,CO_2}) tension); 2) training load; 3) training modality (strength or endurance training) and 4) duration of the training programme (<2 or >2 months of training). The influence of having sham or no sham control group was considered in a separate

analysis. To determine possible association between, on the one hand, methodological quality of individual studies and, on the other hand, individual effect size, a Spearman rank correlation coefficient was calculated. Since small studies will typically show more variability among effect sizes than larger studies and small studies with insignificant effects are less likely published, a funnel plot of sample size *versus* estimated effect size was derived for included studies for *P*_{I,max} (fig. S2 in the online supplementary material). For all outcome variables, the critical value for rejecting H₀ was set two-tailed at p=0.05. On the basis of the classification by Cohen, effect sizes <0.2 were classified as small, those from 0.2 to 0.8 as medium, and those >0.8 as large [24].

RESULTS

Critical review

A total of 129 new studies (time frame 2000-2009) were retrieved from the literature search and 111 studies were excluded (see online supplementary material and fig. S1 for details). This resulted in 18 new eligible studies [25-42] and 14 studies [43-56] published before 2000 from the previous metaanalysis [21]. Accordingly, a total of 32 samples including 830 COPD patients were eligible for meta-analysis. Characteristics of the individual studies are provided in table S3 in the online supplementary material. The overall treatment group consisted of 430 patients (318 males and 112 females) and the control group of 400. The main patient characteristics of the overall treatment group are included in table S4 in the online supplementary material. An overview of the different subgroups studied is provided in table S3 in the online supplementary material. The methodological quality score varied from 30-83% (median 59%) of the maximum score (table S2 in the online supplementary material) and did not significantly relate to the effect size observed in the studies.

Inspiratory muscle strength

The meta-analysis for inspiratory muscle strength demonstrated a significant effect with an increase of 13 cmH₂O (summary effect size (SES) 0.68, 95% CI 0.54–0.82; p<0.001) in the treatment group compared with the control group (table 1 and fig. 1). The homogeneity test, however, showed borderline heterogeneity (I^2 =46%) in the sample, but a random effects model showed similar results (SES 0.73, 95% CI 0.53–0.93; I^2 =37%, p<0.001). As shown in figure 1, the observed heterogeneity between studies was mainly due to a study of DI MAMBRO *et al.* [28]. Subsequent sensitivity analysis without the study of DI MAMBRO *et al.* [28] showed a homogenous SES, which was, again, highly significant (p<0.001). Both studies with sham (n=23) or no sham (n=9) control intervention revealed similar significant improvements in *P*_{L,max}.

Subgroup analysis of inspiratory muscle strength training (k=28) and studies investigating inspiratory muscle endurance training (k=3) revealed that the former remained significant (SES 0.75, 95% CI 0.57–0.87; +13 cmH₂O; I²=45%, p<0.001) whereas the latter lost statistical significance (SES 0.19, 95% CI -0.22–0.59; I²=0%, p=0.36) (fig. 1). No difference was found between threshold and resistive training. The subgroup analysis of GER+IMT *versus* GER showed a homogeneous and significant improvement in inspiratory muscle strength (SES 0.69, 95% CI 0.42–0.96; I²=1%, p<0.001) (fig. 2) in the

| Outcome measures | Subjects n | Q-statistic | l ² | SES | 95% CI | p-value (z-statistic) | Natural units |
|---|------------|-------------|----------------|-------|------------|-----------------------|---|
| PI,max | 32 | 57.8 | 46 | 0.73 | 0.53–0.93 | 0.001 | +13 cmH ₂ O |
| RMET | 14 | 47.3 | 73 | 1.05 | 0.62-1.49 | 0.001 | +261 s |
| ITL | 11 | 16.8 | 3 | 0.98 | 0.72-1.25 | 0.001 | +13 cmH ₂ O |
| MVV | 4 | 1.2 | 0 | 0.23 | -0.27-0.72 | 0.373 | +3 L·min ⁻¹ |
| Functional exercise capacity | 22 | 14.3 | 0 | 0.28 | 0.12-0.44 | 0.001 | 6MWD: +32 m |
| | | | | | | | 12MWD: +85 m |
| Endurance exercise capacity | 3 | 4.6 | 57 | 0.72 | -0.12-1.55 | 0.087 | +198 s |
| V′O₂,max L⋅min ⁻¹ | 9 | 6.0 | 0 | -0.13 | -0.38-0.11 | 0.293 | -0.04 L·min ⁻¹ |
| V′O₂,max mL⋅min ⁻¹ ⋅kg ⁻¹ | 5 | 5.0 | 20 | 0.3 | -0.02-0.63 | 0.067 | +1.3 mL·min ⁻¹ ·kg ⁻¹ |
| V'E,max | 9 | 5.5 | 0 | -0.04 | -0.3-0.2 | 0.696 | -0.7 L·min ⁻¹ |
| Wmax | 10 | 5.1 | 0 | 0.07 | -0.16-0.3 | 0.562 | +1.7 W |
| Dyspnoea Borg score | 14 | 15.6 | 17 | -0.45 | -0.660.24 | 0.001 | -0.9 |
| Dyspnoea TDI | 4 | 6.3 | 52 | 1.58 | 0.86-2.3 | 0.001 | +2.8 |
| Dyspnoea CRQ-Dyspnoea | 9 | 16.6 | 52 | 0.34 | -0.03-0.71 | 0.068 | +1.1 |
| Quality of life CRQ | 9 | 10.4 | 20 | 0.34 | 0.09-0.60 | 0.007 | +3.8 |
| CRQ fatigue | 10 | 8.2 | 0 | 0.27 | 0.03-0.50 | 0.024 | +0.9 |
| CRQ emotion | 10 | 7.6 | 0 | 0.19 | -0.04-0.42 | 0.107 | +0.5 |
| CRQ mastery | 10 | 8.5 | 0 | 0.09 | -0.14-0.33 | 0.432 | -0.005 |

n=32. SES: summary effect size; *P*_{1,max}: maximal inspiratory pressure; RMET: respiratory muscle endurance test: ITL: incremental threshold loading; MVV: maximal voluntary ventilation; *V*'O₂,max: maximal oxygen uptake; *V*'E,max: maximal minute ventilation; Wmax: maximal power output; TDI: transition dyspnoea index; CRQ: chronic respiratory questionnaire.

training group compared with the control group. Further analysis of the studies without inspiratory muscle weakness (k=5) and studies with inspiratory muscle weakness (k=4) demonstrated no significance in the former (SES 0.42, 95% CI -0.01–0.85; I^2 =0%, p=0.06), while the latter remained significant (SES 0.87, 95% CI 0.52–1.21; I^2 =0%, p<0.001).

A significant inverse association was found between baseline P_{a,O_2} and changes in $P_{I,max}$ (r= -0.75, p<0.01; fig. S3, lower panel, in the online supplementary material). In addition, baseline P_{a,CO_2} correlated significantly with changes in $P_{I,max}$ (r=0.61, p<0.05; fig. S3, upper panel, in the online supplementary material), whereas length of the training period and training load did not influence the change in $P_{I,max}$.

Inspiratory muscle endurance

Two meta-analyses for measures of inspiratory muscle endurance were conducted. A significant and homogeneous improvement was found for incremental threshold loading (+13 cmH₂O; SES 0.98, 95% CI 0.72–1.25; I²=3%, p<0.001) in favour of the training group (table 1). Both studies with sham (n=8) or no sham (n=3) control intervention revealed similar significant improvements in endurance load. The analysis for respiratory muscle endurance test (RMET) also produced a significant (+261 s; SES 0.82, 95% CI 0.60–1.04; p<0.001), but heterogeneous (I²=73%) result. After applying a random effects model, significance remained (SES 1.05, 95% CI 0.62–1.49; p<0.001). Both studies with sham (n=9) or no sham (n=5) control intervention revealed similar significant improvements in endurance time.

For subgroup analysis the sample was divided into studies that applied inspiratory muscle strength training (k=11) and

endurance training (k=3). Both samples were heterogeneous $(I^2=52\%$ and 85%); hence, random effects models were applied which demonstrated significant effects in both groups (strength training SES 0.78, 95% CI 0.4-1.15; p<0.0001; endurance training SES 1.95, 95% CI 0.63-3.26; p<0.01). In the strength training group, no difference was found between outcomes for threshold and resistive training devices. In the group with strength training, a significant, homogeneous effect in favour of the IMT group was observed in studies with inspiratory muscle weakness (k=5, SES 0.34, 95% CI 0.04–0.64; $I^2=42\%$, p<0.05). Studies without inspiratory muscle weakness revealed heterogeneous effects (k=6, SES 1.15, 95% CI 0.73-1.57; $I^2=60\%$, p<0.001), while the random model failed to reach statistical significance. The analysis of RMET in studies adding IMT to GER also showed a significant effect in favour of the IMT group (+242 s, SES 0.98, 95% CI 0.72–1.25; I²=3%, p = 0.001).

Exercise capacity

Analysis of functional exercise capacity revealed homogeneous and significant improvement of 6MWD and 12MWD (+32 m and +85 m, respectively) (SES 0.28, 95% CI 0.12–0.44; $I^2=0\%$, p<0.001) in the training group compared with the control group. When comparing inspiratory muscle strength (k=19) and endurance training (k=3), a significant effect was found in favour of the treatment group following strength training (SES 0.29, 95% CI 0.11–0.47; p<0.01), but not for endurance training (SES 0.21, 95% CI -0.19–0.62; p=0.29) (fig. 3). Further analysis showed that the group receiving resistive training (k=3) just lost significance (SES 0.41, 95% CI -0.03–0.85; p=0.06), while the threshold group (k=16) remained significant (SES 0.27, 95% CI 0.07–0.46; p<0.01). Studies with sham control (n=17)

| Study | Year | ÷ | Association measure with 95% Cl |
|--|--------|---------------------------------------|------------------------------------|
| BECKERMAN <i>et al.</i> [25] | 2005 | -8- | 1.39 (1.18–1.60) |
| BELMAN and SHADMEHR [43] | 1988 | | 1.41 (0.88–1.93) |
| BERRY et al. [44] | 1996 — | | -0.04 (-0.50-0.43) |
| BUSTAMANTE MADARIAGA et al. [26] | 2007 | | 0.84 (0.47-1.22) |
| Covey <i>et al.</i> [27] | 2001 | | 0.78 (0.48-1.09) |
| Dekhuijzen <i>et al</i> . [45] | 1991 | | 0.57 (0.37–0.77) |
| Di Мамвко <i>et al</i> . [28] [#] | 2007 | | 4.50 (3.73–5.26) |
| GOLDSTEIN et al. [46] | 1989 — | | -0.04 (-0.76–0.68) |
| HARVER et al. [47] | 1989 | | 0.53 (0.11–0.95) |
| Heijdra <i>et al.</i> [48] | 1996 | | 1.97 (1.49–2.46) |
| HILL <i>et al.</i> [29] | 2006 | | 0.70 (0.45–0.95) |
| Hsiao <i>et al</i> . [30] | 2003 | _= | 0.87 (0.46–1.28) |
| Hsiao <i>et al</i> . [30] | 2003 | | 0.55 (0.15–0.95) |
| Кім <i>et al</i> . [49] | 1993 | - | 0.20 (0.07-0.32) |
| Larson <i>et al</i> . [50] | 1988 | | 0.29 (-0.09–0.67) |
| Larson <i>et al</i> . [32] | 1999 | | 0.40 (0.08–0.72) |
| Larson <i>et al</i> . [32] | 1999 | | 0.44 (0.16–0.73) |
| LISBOA <i>et al</i> . [51] | 1997 | | 0.71 (0.31–1.12) |
| MAGADLE et al. [34] | 2007 | | 0.84 (0.57–1.10) |
| PATESSIO <i>et al.</i> [52] | 1989 | | 0.47 (-0.02–0.97) |
| PREUSSER <i>et al.</i> [53] | 1994 | | 0.43 (0.02–0.85) |
| RAMIREZ-SARMIENTO et al. [35] | 2002 | | 1.11 (0.51–1.72) |
| SÁNCHEZ RIERA et al. [36] | 2001 | | 1.62 (1.16–2.08) |
| Serón <i>et al</i> . [38] | 2005 | | 0.35 (0.13–0.58) |
| Sykes [39] | 2005 | | 0.87 (0.65–1.10) |
| VILLAFRANCA <i>et al.</i> [54] | 1998 | | 0.69 (0.29–1.10) |
| WANKE <i>et al.</i> [55] | 1994 | + | 0.97 (0.77–1.16) |
| WEINER <i>et al</i> . [56] | 1992 | | 1.27 (0.91–1.63) |
| WEINER et al. [41] | 2003 | | 1.88 (1.28–2.48) |
| SES inspiratory muscle strength traini | ng | \diamond | 0.72 (0.57–0.87) |
| KOPPERS <i>et al.</i> [31] | 2006 | | 0.15 (-0.07–0.37) |
| Mador <i>et al.</i> [33] | 2005 | | 0.19 (-0.08–0.47) |
| SCHERER et al. [37] | 2000 | | 0.23 (-0.04–0.49) |
| SES inspiratory muscle endurance tra | ining | \diamond | 0.19 (-0.22–0.59) |
| Total SES | | \diamond | 0.65 (0.51–0.79) |
| | [] | · · · · · · · · · · · · · · · · · · · | . , |
| | -2 -1 | 0 1 2 3 4 | |
| | | PI,max cmH ₂ O | |

FIGURE 1. Effect of inspiratory muscle strength and endurance training on maximal inspiratory strength (PI,max). #: excluded from the analyses.

revealed significant improvements in functional exercise performance (SES 0.30, 95% CI 0.11–0.49; $I^2=0\%$, p<0.001), but studies without sham control intervention (n=5) failed to reach statistical significance (p=0.09). The subgroup analysis of studies applying IMT as a mono-intervention found a significant improvement in favour of the IMT group (SES 0.27, 95% CI 0.07–0.46; p<0.01), while studies adding IMT to GER (k=7) were marginally significant (SES 0.29, 95% CI -0.004–0.58; p=0.05). Subdividing this analysis into a group of studies with (k=3) and a group without (k=4) inspiratory muscle weakness, resulted in a significant result for the former (SES

0.51, 95% CI 0.10–0.91; p<0.05), but not for the latter (SES 0.05, 95% CI -0.37–0.47; p=0.82) (fig. 4). However, the muscle weakness subgroup showed heterogeneity and just failed to reach statistical significance (SES 0.60, 95% CI -0.08–1.29; p=0.08) in the random effects model. The analysis of endurance exercise capacity produced a significant (k=3, SES 0.80, 95% CI 0.32–1.37; +198 s; p<0.01), though not homogeneous (I²=57%), improvement in the IMT group compared with the control group. The subsequent random effects model did not reach statistical significance (SES 0.72, 95% CI -0.12–1.55; p=0.08).

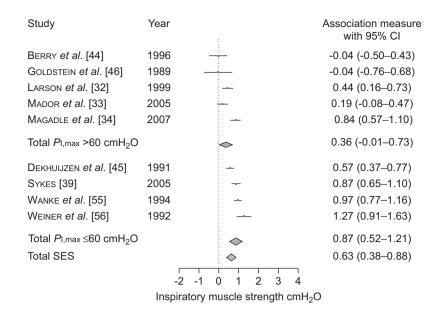


FIGURE 2. Effect of general exercise reconditioning (GER) plus inspiratory muscle training versus GER alone on inspiratory muscle strength. Pl,max: maximal inspiratory pressure.

Results for maximal power output (Wmax) on the bicycle test (k=10, SES 0.07, 95% CI -0.16–0.30; +2 W; I²=0%, p=0.56), maximal minute ventilation (V'E,max) (k=9, SES -0.04, 95% CI -0.30–0.21; -0.7 L·min⁻¹; I²=0%, p=0.69), maximal oxygen uptake (V'O₂,max) in L·min⁻¹ (k=9, SES -0.13, 95% CI -0.38–0.11; 0.04 L·min⁻¹; I²=0%, p=0.29) and oxygen uptake in mL·min⁻¹·kg⁻¹ (k=5, SES 0.30, 95% CI -0.02–0.63; 1.3 mL·min⁻¹·kg⁻¹; I²=20%, p=0.06) revealed no significant overall effect. Subgroup analyses for Wmax, V'E,max and V'O₂,max revealed no significant effect for either of the subgroups IMT *versus* control and IMT+GER *versus* GER. The analysis conducted for maximal voluntary ventilation (MVV) did show a homogenous (I²=0%) and nonsignificant summary effect (SES 0.23, 95% CI -0.27–0.72; +3 L·min⁻¹; p=0.37).

Dyspnoea

Three analyses containing measures of dyspnoea (Borg score, transition dyspnoea index (TDI) and chronic respiratory questionnaire (CRQ) dyspnoea score) were conducted. A significant and homogeneous (I²=17%) effect was found for Borg score (SES -0.45, 95% CI -0.66– -0.24; -0.9 point; p<0.0001), favouring the training group. When comparing threshold and resistive training for this measure, the former remained significant (SES -0.38, 95% CI -0.62- -0.14; p<0.01) while the latter lost significance (SES -0.46, 95% CI -0.99-0.07; p=0.09). Studies with sham (n=8) revealed significant improvements in dyspnoea (SES -0.67, 95% CI -0.96– -0.40; p<0.0001), but studies without sham control intervention (n=6) were not statistically significant (p=0.16). The analysis conducted for the focal score of the TDI demonstrated a significant, but heterogeneous (I²=52%), effect (SES 1.47, 95% CI 0.99–1.96; 2.8 points; p<0.001) in favour of the treatment group. The result of the random effects model remained significant (SES 1.58, 95% CI 0.86–2.30; p<0.0001). Similarly, the analysis conducted for CRQ dyspnoea score showed a significant, but heterogeneous (I²=52%), increase (SES 0.33, 95% CI 0.07–0.58; 1.1 points;

p<0.05). The random effects model failed to show statistical significance (SES 0.34, 95% CI -0.03–0.71; p=0.06). Subgroup analysis for CRQ dyspnoea in studies that applied inspiratory muscle strength training (k=8) demonstrated a significant, homogeneous effect (SES 0.45, 95% CI 0.18–0.72; p<0.01) in favour of the IMT group. Studies with sham (n=4) revealed significant improvements in CRQ dyspnoea (SES 0.44, 95% CI 0.06–0.82; p<0.01), but studies without sham (n=5) control intervention failed to reach statistical significance (p=0.08). Subgroup analysis for endurance training was not possible, due to the limited number of studies using this outcome (k=1). A subgroup analysis of trials adding IMT to a GER programme showed no significant effects on the Borg dyspnoea score (SES -0.32, 95% CI -0.70–0.06; p=0.10) and CRQ dyspnoea score (SES -0.06, 95% CI -0.78–0.66; p=0.88).

Quality of life

Meta-analysis demonstrated a significant (SES 0.34, 95% CI 0.09–0.60; 3.8 points; p<0.01), homogeneous (I²=23%) improvement in the total score of the CRQ in favour of the IMT group when compared with the control group. Studies with sham (n=5) revealed significant improvements in quality of life (SES 0.40, 95% CI 0.07–0.73; p<0.01), but studies without sham (n=4) control intervention failed to reach statistical significance (p=0.08). Analysis for the different items of the CRQ scale demonstrated a significant effect on the dyspnoea (see above) and the fatigue scale (+0.87 points; p < 0.05) in favour of the IMT group. The analysis for emotion failed to show statistical significance (SES 0.19, 95% CI -0.04-0.42; +0.48 points; p=0.10), while mastery showed no improvement (SES 0.09, 95% CI -014-0.33; -0.01 points; p=0.42). No effects on these items were found when performing the subgroup analysis of IMT+GER versus GER. Meta-analyses could not be performed on other measurements of quality of life (see online supplementary material).

| Study | Year | | Association measure with 95% Cl |
|----------------------------------|--------------|----------------------------|------------------------------------|
| BECKERMAN et al. [25] | 2005 | - | 0.46 (0.27-0.64) |
| BERRY et al. [44] | 1996 | | -0.01 (-0.47–0.46) |
| DEKHUIJZEN <i>et al.</i> [45] | 1991 | - | 0.22 (0.03-0.42) |
| Di Mambro <i>et al</i> . [28] | 2007 | <u>_</u> | -0.11 (-0.51–0.28) |
| GOLDSTEIN et al. [46] | 1989 | | -0.03 (-0.75–0.69) |
| HILL <i>et al</i> . [29] | 2006 | ÷ | 0.21 (-0.06-0.49) |
| Hsiao <i>et al</i> . [30] | 2003 | | 0.30 (-0.10-0.69) |
| Hsiao <i>et al</i> . [30] | 2003 | | 0.30 (-0.09–0.70) |
| Кім <i>et al.</i> [49] | 1993 | | 0.07 (-0.07-0.21) |
| Larson <i>et al</i> . [50] | 1988 | : | 0.32 (-0.09–0.73) |
| LISBOA <i>et al</i> . [51] | 1997 | | 0.66 (0.26-1.07) |
| MAGADLE et al. [34] | 2007 | | 0.19 (-0.07–0.44) |
| PREUSSER et al. [53] | 1994 | | 0.11 (-0.30-0.52) |
| RAMIREZ-SARMIENTO et al. [35] | 2002 | | 0.10 (-0.46–0.66) |
| SÁNCHEZ RIERA <i>et al.</i> [36] | 2001 | | 0.97 (0.55–1.38) |
| Serón <i>et al.</i> [38] | 2005 | | -0.04 (-0.26-0.19) |
| Sykes [39] | 2005 | | 0.33 (0.11–0.54) |
| WEINER <i>et al.</i> [56] | 1992 | | 1.48 (1.11–1.86) |
| WEINER et al. [41] | 2003 | | 0.63 (0.13–1.14) |
| SES inspiratory muscle strengt | h training | • | 0.29 (0.11–0.47) |
| KOPPERS et al. [31] | 2006 | - | 0.34 (0.12-0.56) |
| Mador <i>et al.</i> [33] | 2005 | | -0.04 (-0.31-0.23) |
| SCHERER et al. [37] | 2000 | | 0.32 (0.06–0.59) |
| SES inspiratory muscle endura | nce training | ~ | 0.21 (-0.19–0.62) |
| Total SES | | ♦ | 0.28 (0.12-0.44) |
| | -2 | -1 0 1 2 3 4 | |
| | Fu | nctional exercise capacity | |

FIGURE 3. Effect of inspiratory strength and endurance training on functional exercise capacity (6-min walk distance).

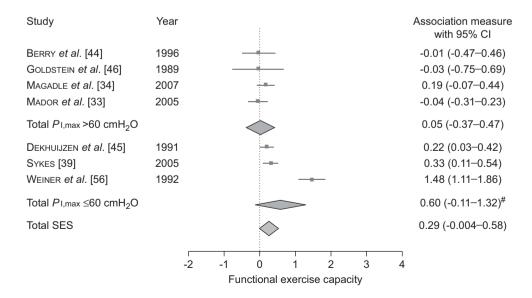


FIGURE 4. Effect of general exercise reconditioning (GER) plus inspiratory muscle training versus GER alone on functional exercise performance. Pl,max: maximal inspiratory pressure. *: result from random effects model.

DISCUSSION

This meta-analysis included 18 more studies and further substantiates the evidence for application of IMT in COPD patients. Statistically significant medium to large effect sizes (mean 0.73) and clinical relevant improvements were observed for inspiratory muscle strength (13 cmH₂O) and endurance (261 s), functional exercise capacity (32 m), dyspnoea (2.8 points TDI) and quality of life (3.8 points CRQ). Changes in maximal and endurance exercise capacity failed to reach statistical significance. Inspiratory muscle strength training was superior to endurance training for improving inspiratory muscle strength, functional exercise performance and dyspnoea. Patients with a clear inspiratory muscle weakness and lower Pa,O2 or higher Pa,CO2 showed better response to IMT on inspiratory muscle strength and functional exercise capacity than those patients with better preserved respiratory muscle function. The improvement in inspiratory muscle endurance capacity is supported by a study in COPD patients showing significant increases in the proportion of type I fibres and size of type II fibres in the external intercostals after IMT [35]. The addition of IMT to a general exercise training programme resulted, only in patients with inspiratory muscle weakness, in borderline nonsignificant improvement of functional exercise performance.

Although we agree that an appropriately designed randomised controlled trial should be the golden standard to test the hypothesis that IMT is effective in patients with COPD, we should also acknowledge that most randomised controlled trials suffer from insufficient statistical power to underpin the effects of IMT. With that, the strength of a meta-analysis is the ability to combine studies with the same intervention and same outcome "as if" there is one overall study in order to prevent type II error. Therefore, conclusions from meta-analyses are more powerful than from a single randomised controlled trial. A second strength of meta-analysis is the ability to compare individual studies, allowing the identification of possible modifiers and confounders that may have affected claimed effects in order to give recommendations for further research.

Our meta-analysis confirms and debates several findings of the recent meta-analysis by GEDDES et al. [57]. In addition, we provide new data on the choice of training modalities for IMT and patient characteristics relevant for selection of patients. The analysis by GEDDES et al. [57] contained substantially fewer papers than the present meta-analysis. For 6MWD, maximal exercise capacity, dyspnoea Borg and quality of life, their respective analyses contained four, four, four and only two trials, while we reported on 22, 10, 14 and nine papers, respectively. This is, in part, due to inclusion of studies in other languages (n=2) [26, 28], studies that added IMT to general exercise training (n=8) [33, 34, 39, 40, 44, 45, 55, 56], recent publication (n=3) [26, 28, 34], and endurance training modality (n=2) [33, 37]. In addition, we included studies with and without a "sham training" control group. Statistical significance was lost for the analysis (including a very small number of studies) with comparison of IMT with no sham intervention for functional exercise capacity, dyspnoea and quality of life. However, the meta-analysis performed in studies with comparison of IMT with a sham control intervention, the most critical comparison, revealed statistical significant improvements in

inspiratory muscle strength and endurance, functional exercise capacity, dyspnoea and quality of life.

The analysis of functional exercise capacity included trials measuring 6MWD, 12MWD and one paper that used the shuttle walk test (SWT) [36]. The analysis showed no heterogeneity and analyses of functional exercise capacity without the data from the SWT did not affect any of the results. The clinical importance of the findings was investigated through transforming the outcomes to natural units. The improvement of 32 m on the 6MWD was highly significant and within the range of the recently redefined minimal clinical important difference (MCID) of 35 m [58] and 25 m [59]. The effects on quality of life (CRQ +3.8 points) and its sub-items did not reach the MCID of 0.5 points per question. The TDI score of 2.8 points clearly crosses the MCID of 1 point [60].

Previous meta-analyses [20, 57] reported significant effects on maximal ventilation during exercise and MVV that were not observed in our analyses. However, their analysis on MVV included two reports [43, 47], while ours included two extra trials [44, 46]. Similarly, their analysis on *V*'E_{,max} included only two reports [36, 51], while our analysis was based on nine reports [29, 32, 33, 36, 44, 45, 51, 55]. It is likely that their limited number of studies caused false-positive effects.

Does the choice for inspiratory muscle strength or endurance training matter?

Three types of inspiratory muscle training, i.e. "inspiratory resistive training", "threshold loading" and "normocapnic hyperpnoea" (NCH) were practised. NCH is considered "endurance training" since it has high number of repetitions and no additional resistance for 15-20 min [61, 62]. "Inspiratory resistive breathing" and "threshold loading" are considered a mixture of strength and endurance training, since the number of contractions is low and the inspiratory resistance is relatively high. Subgroup analyses comparing strength and endurance training revealed better effects on inspiratory muscle strength, functional exercise capacity and dyspnoea following strength training. Both modalities of training significantly improved inspiratory muscle endurance, but only strength training was able to significantly improve PI,max and functional exercise capacity. It is, however, difficult to draw firm conclusions, since only three studies investigated the effects of inspiratory muscle endurance training, while 29 trials applied strength training. A head to head comparison between strength and endurance training is needed to substantiate differences in effectiveness.

Subgroup analysis of resistive breathing and threshold training for functional exercise capacity and dyspnoea measured on the Borg scale revealed no statistical significant effect for the group receiving resistive breathing. BELMAN *et al.* [63] showed that similar workloads were obtained during resistive breathing and threshold loading. Threshold loading, however, enhances velocity of inspiratory muscle shortening and thus will shorten inspiratory time [54]. This allows more time for exhalation and might reduce dynamic hyperinflation [64]. The increased relaxation time might prevent the development of inspiratory muscle fatigue [65]. Furthermore, threshold loading significantly increased peak inspiratory flow rate [42]. This might have clinical relevance for patients with severe COPD who are not able to generate adequate flow to secure optimal lung deposition of inhaled medication. Finally, inspiratory pressure during resistive loading is flow dependent [43]. Since this is not the case for threshold loading [66], the latter device is preferred, especially for unsupervised training at home.

Training intensity varied among studies from 30% of *P*_{I,max} [50, 51, 53] up to 50–80% of *P*_{I,max} [45, 53, 55]. We did not observe a dose–response relationship, probably because inclusion of studies in the meta-analysis was set at an inspiratory load of \geq 30% *P*_{I,max}.

Does IMT add to a general exercise training programme?

IMT, when added to a GER programme, had an additional effect on inspiratory muscle strength and endurance, but not on dyspnoea (Borg score and CRQ scale). Similar to the effects observed during IMT alone, no additional effect was found on measures of maximal exercise capacity (Wmax, V'E,max and $V'O_{2,max}$). These conclusions are at variance with the analysis of O'BRIEN et al. [20], who observed a significant increase in maximal exercise tidal volume, but not on inspiratory muscle endurance. In the present meta-analysis, the additional effect of IMT to GER on functional exercise capacity just failed to reach statistical significance in patients with inspiratory muscle weakness. However, the observed strong trend in the present analysis, including two more studies, was stronger than in our previous meta-analysis [21]. It has face validity to consider that patients limited in their exercise performance by ventilatory constraints and dyspnoea due to inspiratory muscle weakness might indeed benefit from additional IMT. However, this hypothesis has to be investigated in a randomised controlled trial. Furthermore, inspiratory muscle fatigue reflexively induces sympathetically mediated vasoconstrictor activity, thereby compromising blood flow to the active limb muscles [67]. In turn, blood flow and oxygen transport to the working muscle are reduced, thereby exacerbating limb fatigue and compromising exercise performance [68]. Further research has to provide evidence that IMT indeed reduces the onset of respiratory muscle "fatigue", and thereby improves exercise performance.

Which COPD patient will benefit from IMT?

In the present analysis we observed that patients with inspiratory muscle weakness, defined as PI,max <60 cmH₂O, were more likely to improve inspiratory muscle strength and functional exercise performance when IMT was applied in addition to GER. In addition, studies including patients with a low baseline Pa,O2 or high baseline Pa,CO2 had more improvement in inspiratory muscle strength achieved by IMT (fig. S3 in the online supplementary material). Indeed, inspiratory muscle weakness is associated with hypercapnia [3] and nocturnal desaturation [6], while IMT improved nocturnal desaturation significantly [48]. Improvement of dyspnoea was a persistent finding in the present study. The increased relative work of breathing is indicated by the high PI/PI,max ratio, which is related to the perceived dyspnoea sensation [9]. IMT will increase PI,max, reduce subsequently the PI/PI,max ratio and thus dyspnoea sensation. Therefore, especially patients with high levels of dyspnoea might benefit from IMT.

Since most studies show loss of training effects on the long term, patients should be encouraged to continue with the IMT.

One study showed that the long-term benefits were maintained only when the IMT was continued [69].

Conclusion

In conclusion, IMT is an effective treatment modality in COPD patients to improve respiratory muscle strength and endurance, resulting in reductions of dyspnoea and improvement in functional exercise capacity and HRQoL. Patients with more advanced muscle weakness seem to be better responders, especially when considering IMT in addition to general exercise training. Inspiratory muscle endurance training was shown to be less effective than respiratory muscle strength training.

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STATEMENT OF INTEREST

A statement of interest for M. Decramer can be found at www.erj. ersjournals.com/site/misc/statements.xhtml

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